

STATISTICS

Module 2 of MATHEMATICS AND STATISTICS

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L5. Conditional probability and Bayes formula.

Outline

Conditional Probabilities

- Definition

- The Multiplication Rule

- Bayes Formula and Partition rule

Introduction

Conditional probabilities are useful

- when we update the probability of events according to the information;
- easy way for computing probabilities of complex events.

Example: Two (green and red) dice are rolled. The sample space is $S = \{(i, j); i, j : 1, \dots, 6\}$. Consider the event $E = \{\text{sum is 8}\}$.

$$\Rightarrow \mathbb{P}(E) = \frac{\#(E)}{\#(S)} = \frac{5}{36}.$$

If the event **red die shows 6** has occurred, how will the probability of E be modified?

Let $F = \{\text{red die is 6}\}$ and denote by $\mathbb{P}(E|F)$ the “*conditional probability of E given F* ”

Intuition: If F has occurred, the event E can occur if and only if $E \cap F$ occurs. Equivalently, $\mathbb{P}(E|F)$ will be the the probability that (6,2) occurs out of the 6 outcomes of $F = \{(6, 1), (6, 2), (6, 3), (6, 4), (6, 5), (6, 6)\}$.

$$\Rightarrow \mathbb{P}(E|F) = \frac{\#(E \cap F)}{\#(F)} = \frac{1}{6}.$$

Conditional Probability

Def. Consider two events E and F with $\mathbb{P}(F) > 0$. The conditional probability of E given F is

$$\mathbb{P}(E|F) = \frac{\mathbb{P}(E \cap F)}{\mathbb{P}(F)}.$$

Example: A fair coin is flipped twice. Compute the probability of $E=\{2 \text{ heads}\}$ conditionally to

- a) $F=\{\text{head at first flip}\}$
- b) $G=\{\text{at least one head}\}$.

▷ $S = \{(H, H), (T, H), (H, T), (T, T)\}$ are the **equally likely** 4 possible outcomes.

$E = \{(H, H)\}$, $F = \{(H, T), (H, H)\}$ e $G = \{(H, H), (T, H), (H, T)\}$.

$$\Rightarrow \text{a) } \mathbb{P}(E|F) = \frac{\mathbb{P}(E \cap F)}{\mathbb{P}(F)} = \frac{\mathbb{P}((H, H))}{\mathbb{P}(\{(H, H), (H, T)\})} = \frac{\#((H, H))}{\#(\{(H, H), (H, T)\})} = \frac{1}{2}$$

$$\Rightarrow \text{b) } \mathbb{P}(E|G) = \frac{\mathbb{P}(E \cap G)}{\mathbb{P}(G)} = \frac{\mathbb{P}((H, H))}{\mathbb{P}(\{(H, H), (H, T), (T, H)\})} = \frac{1}{3}$$

By the definition we immediately have

$$\mathbb{P}(E \cap F) = \mathbb{P}(F)\mathbb{P}(E|F),$$

which is used to compute the probability of the intersection.

Let us generalize to n events.

The Multiplication Rule

Prop. Let $\mathbb{P}(E_1 \cap E_2 \cap \dots E_{n-1}) > 0$, then

$$\mathbb{P}(E_1 \cap E_2 \cap \dots E_n) = \mathbb{P}(E_1) \mathbb{P}(E_2|E_1) \mathbb{P}(E_3|E_1 \cap E_2) \cdots \mathbb{P}(E_n|E_1 \cap E_2 \dots E_{n-1}).$$

Proof follows from the definition of conditional probability. Note that $\mathbb{P}(E_1 \cap E_2 \cap \dots E_{n-1}) > 0$ is sufficient for all conditional probabilities appearing in the formula to be well defined.

Example: An urn contains 8 red and 4 white balls. Three balls are drawn without replacement, what is the probability $E = \{2 \text{ red and } 1 \text{ white}\}$ occurs? Denote by R_i (W_i) “red (white) at the i -th draw”.

$\mathbb{P}(E) = \mathbb{P}((W_1 R_2 R_3) \cup (R_1 W_2 R_3) \cup (R_1 R_2 W_3))$ given the three events are mutually exclusive,

$$\mathbb{P}(E) = \mathbb{P}(W_1 R_2 R_3) + \mathbb{P}(R_1 W_2 R_3) + \mathbb{P}(R_1 R_2 W_3).$$

$$\mathbb{P}(W_1 R_2 R_3) = \mathbb{P}(W_1) \mathbb{P}(R_2|W_1) \mathbb{P}(R_3|W_1 R_2) = \frac{4}{12} \frac{8}{11} \frac{7}{10} = \frac{28}{165}$$

Besides, the three events have the same probability $\Rightarrow$

$$\mathbb{P}(E) = 3 \frac{28}{165} = \frac{28}{55}$$

$\mathbb{P}(\cdot|F)$ is a probability measure.

Prop. Given a sample space S , a family of events, a probability $\mathbb{P}$ and an event F with $\mathbb{P}(F) > 0$, the conditional probability $\mathbb{P}(\cdot|F)$ is a *probability measure* on the sample space. S .

▷ Let us check $\mathbb{P}(\cdot|F)$ satisfies the axioms:

A.1 for any E , $0 \leq \mathbb{P}(E|F) = \frac{\mathbb{P}(E \cap F)}{\mathbb{P}(F)} \leq 1$

A.2 $\mathbb{P}(S|F) = \frac{\mathbb{P}(S \cap F)}{\mathbb{P}(F)} = \frac{\mathbb{P}(F)}{\mathbb{P}(F)} = 1$

A.3 If $\{E_n, n \geq 1\}$ is a sequence of mutually exclusive events, then

$$\mathbb{P}\left(\bigcup_{n:1}^{\infty} E_n | F\right) = \sum_{n:1}^{\infty} \mathbb{P}(E_n | F).$$

By the definition of conditional probability,

$$\mathbb{P}\left(\bigcup_{n:1}^{\infty} E_n | F\right) = \frac{\mathbb{P}\left(\left(\bigcup_{n:1}^{\infty} E_n\right) \cap F\right)}{\mathbb{P}(F)}.$$

By the distributive law and axiom A.3 of $\mathbb{P}$ (note that $(E_n \cap F) \cap (E_m \cap F) = \emptyset$ if $n \neq m$), we compute the numerator

$$\mathbb{P}\left(\bigcup_{n:1}^{\infty} (E_n \cap F)\right) = \sum_{n:1}^{\infty} \mathbb{P}(E_n \cap F).$$

Dividing by $\mathbb{P}(F)$, the result follows.

Partition rule

We have for any events E, F , with $0 < \mathbb{P}(F) < 1$

$$\mathbb{P}(E) = \mathbb{P}(F)\mathbb{P}(E|F) + \mathbb{P}(F^c)\mathbb{P}(E|F^c).$$

- ▷ We can write E as union of two mutually exclusive events

$$E = E \cap (F \cup F^c) = (E \cap F) \cup (E \cap F^c)$$

$$\mathbb{P}(E) = \mathbb{P}(E \cap F) + \mathbb{P}(E \cap F^c) = \mathbb{P}(F)\mathbb{P}(E|F) + \mathbb{P}(F^c)\mathbb{P}(E|F^c).$$

This can be generalized to the following case.

Prop. Consider the events $E, F_1, \dots, F_n$, with $\mathbb{P}(F_i) > 0$ and $\cup_{i=1}^n F_i = S$ e $F_i \cap F_j = \emptyset$ for any $i \neq j$. Then,

$$\mathbb{P}(E) = \sum_{i=1}^n \mathbb{P}(F_i)\mathbb{P}(E|F_i),$$

(*weighted average of $\mathbb{P}(E|F_i)$ with weights $\mathbb{P}(F_i)$*).

The well known **Bayes formula** is byproduct of the *partition rule*.

Bayes Formula

Prop. Consider a partition of S , $F_1, \dots, F_n$, with $\mathbb{P}(F_i) > 0$ and $\cup_{i=1}^n F_i = S$ e $F_i \cap F_j = \emptyset$ for any $i \neq j$ and an event E , $\mathbb{P}(E) > 0$. Then, for any F_j , $j \in \{1, \dots, n\}$

$$\mathbb{P}(F_j|E) = \frac{\mathbb{P}(F_j)\mathbb{P}(E|F_j)}{\sum_{i=1}^n \mathbb{P}(F_i)\mathbb{P}(E|F_i)}$$

- the F_j s stand for alternative “hypothesis”,
- $\mathbb{P}(F_j)$ are known as *priors*,
- $\mathbb{P}(F_j|E)$ the *a posteriori probability* (given E).

Example: An insurance company believes that people can be divided into 2 classes: those who are accident prone and those who are not. Their statistics shows that an accident-prone person will have an accident within a year with probability 0.4, while this probability is 0.2 for a person belonging to the other class. If the accident prone persons are the 30% of the population, what is the probability that a new policyholder will have an accident within a year? If he had, what is the probability that he is an accident prone person?

Monty Hall Problem: Goats and cars

Three doors A,B,C are closed. Behind one of them there is a Ferrari, whereas the others hide two goats. The player selects one door, say A for sake of simplicity. The TV presenter, who knows where the car is, tells her/him that the car is not behind door B and gives him the possibility to revise her/his choice (i.e. that is to select door C). Should the player revise her/his choice?